

Uploaded to VFC Website

~ October 2012 ~

This Document has been provided to you courtesy of Veterans-For-Change!

Feel free to pass to any veteran who might be able to use this information!

For thousands more files like this and hundreds of links to useful information, and hundreds of "Frequently Asked Questions, please go to:

Veterans-For-Change

Veterans-For-Change is a 501(c)(3) Non-Profit Corporation Tax ID #27-3820181

If Veteran's don't help Veteran's, who will?

We appreciate all donations to continue to provide information and services to Veterans and their families.

https://www.paypal.com/cgi-bin/webscr?cmd= s-xclick&hosted button id=WGT2M5UTB9A78

Note:

VFC is not liable for source information in this document, it is merely provided as a courtesy to our members.

Item D Number 01580

Author Wolfe, William H.

Corporate Author United States Air Force School of Aerospace Medicine (

Report/Article Title An Epidemiologic Investigation of Health Effects in Air

Force Personnel Following Exposure to Herbicides:

Mortality Update - 1985

Journal/Book Title

Year 1985

Menth/Day November 29

Color

Number of Images 49

Descripton Notes

PROJECT RANCH HAND II AN EPIDEMIOLOGIC INVESTIGATION OF HEALTH EFFECTS IN AIR FORCE PERSONNEL FOLLOWING EXPOSURE TO HERBICIDES

MORTALITY UPDATE - 1985 29 NOVEMBER 1985

WILLIAM H. WOLFE, COLONEL, USAF, MC JOEL E.MICHALEK, PhD

Prepared for:

THE SURGEON GENERAL UNITED STATES AIR FORCE WASHINGTON, D.C. 20314



EPIDEMIOLOGY DIVISION DATA SCIENCES DIVISION

USAF SCHOOL OF AEROSPACE MEDICINE (AFSC)
BROOKS AIR FORCE BASE, TEXAS 78235

- ·	REPORT DOCU	MENTATION	PAGE		
18 REPORT SECURITY CLASSIFICATION UNCLASSIFIED		16. RESTRICTIVE	MARKINGS	· · · · · · · · · · · · · · · · · · ·	
28. SECURITY CLASSIFICATION AUTHORITY		3. DISTRIBUTION	I/AVAILABILITY (OF REPORT	
26. DECLASSIFICATION / DOWNGRADING SCHED		for public ion is unli			
PERFORMING ORGANIZATION REPORT NUMBER			ORGANIZATION (R(S)
60. NAME OF PERFORMING ORGANIZATION	6b. OFFICE SYMBOL	7a. NAME OF M	ONITORING ORGA	ANIZATION	
USAF School of Aerospace Medicine	(If applicable) USAFSAM/EK	The Surge	on General		
ADDRESS (City, State, and ZIP Code) Aerospace Medical Division (AF Brooks Air Force Base, Texas		United St	y, State, and ZIP ates Air For n., D.C. 20	rce	
a. NAME OF FUNDING/SPONSORING ORGANIZATION Aerospace Medicine	8b. OFFICE SYMBOL (If applicable) USAFSAM/EK	9. PROCUREMENT	INSTRUMENT ID	ENTIFICATION N	UMBER
C ADDRESS (City, State, and ZIP Code)	^	10. SOURCE OF F	UNDING NUMBER		
Aerospace Medical Division (AF	_		PROJECT NO.	TASK NO.	WORK UNIT ACCESSION NO
Brooks Air Force Base, Texas TITLE (Include Security Classification) An Ep		65306F	2767	00	01
المكافلة المسمود المنادات الأراث ليدي التهاج المتعلق المتعلق المتعلق المستور المتعارف والمناز والمناف والمتعارف	79 701984_]	1985 Novem	RT (Year, Month, i ber 29	Day) 15. PAGE 47	
COSATI CODES FIELD GROUP SUB-GROUP 06 05	18. SUBJECT TERMS (C Epidemiologic I Matched Cohort	1985 Novem ontinue on reverse nvestigation Design; Nonc	if necessary and ; Air Force	identify by bloo	sk number) idy;
6. SUPPLEMENTARY NOTATION 7. COSATI CODES FIELD GROUP SUB-GROUP 06 05	18. SUBJECT TERMS (C Epidemiologic I Matched Cohort Mortality Study and identify by block as salth Study is t loides in Viets is a result of released in Ju collow-up mortal at could be con- leased be con- leased by the th subjects were de- ots had died.	1985 Novementine on reverse nvestigation Design; Noncomber) to determine nam during their particular their particular study in clusively at ird mortality study in termined. The path cortain of the cort	# necessary and; Air Force oncurrent Possessine December tributed to y analyses. In groups. It relative	identify by bion Health Stu rospective dividuals in Hand operathat programme that programme in herbioide Deaths in herbioide	nvolved in ation have gram. The Report in ther study for dioxin n the 1257, 55 Ranch ed on all ignificant mortality d healthy.
COSATI CODES FIELD GROUP SUB-GROUP OB OB ABSTRACT (Continue on reverse if necessary and the purpose of the Air force He can be acrial spraying of herbit experienced increased deaths a seline Mortality Report was rebruary 1984, and the first force the exposure. The present report desired and and 5171 comparison and and 285 comparison subjects. The current mortality between the sufficient time may have obstructed in the sufficient time may have obstructed and another time may have obstructed.	18. SUBJECT TERMS (C Epidemiologic I Matched Cohort Mortality Study and identify by block me salth Study is t loides in Viets as a result of released in Ju collow-up mortal at could be con- leased by mortal at could by mortal at could by mortal at could by mortal at could be con- leased by mortal at could by mortal at could be con- leased by mortal at could be con- leased by mortal at could be con- leased by mortal at could by mortal at could be con- con- at could be con- at	1985 Novement on reverse investigation Design; Noncomber) of determine their particular	if necessary and; Air Force oncurrent P. whether inches Ranch dipation in a Baseline December tributed to y analyses. I ficates we have station groups. I relatively signifully	identify by bion Health Sturospective dividuals in Hand operation that prose Morbidity 1984. Weil herbioide Deaths dividually sember 1984 we obtain stically sember touch Took	ck number) idy; Design; nvolved in ation have gram. The Report in ther study for dioxin in the 1257; 55 Ranch ed on all ignificant mortality di healthy.
COSATI CODES FIELD GROUP SUB-GROUP OB OB ABSTRACT (Continue on reverse if necessary at the purpose of the Air force He the aerial spraying of herbitexperienced increased deaths a sealine Mortality Report was rebruary 1984, and the first filemonstrated health effects the exposure. The present report deaneh Hand and 6171 comparison land and 285 comparison subjects. The current mortality between the entitle sufficient time may have constructed in mortality between the entitle sufficient time may have a sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the comparison of the entitle sufficient time may have constructed as the constructed sufficient time shall be constructed as the constructe	18. SUBJECT TERMS (C) Epidemiologic I Matched Cohort Mortality Study and identify by block acts salth Study is t loides in Viets as a result of released in Ju collow-up mortal at could be con- iescribes the th subjects were de ots had died. Ity analyses die on the study grove elepsed for a	1985 Novementinue on reverse nvestigation Design; Noncomber) to determine during their particular study in clusively at ird mortality study in the s	if necessary and; Air Force oncurrent Possible Ranch dipation in a Baseline December tributed to y analyses. In of 31 December	identify by block Health Sturospective dividuals in Hand operation that programmed that programmed herbioide Deaths in herbioi	ok number) idy; Design; nvolved in ation have gram. The Report in ther study or dioxin n the 1257; 55 Ranch ed on all ignificant mortality d healthy. ditions to
COSATI CODES FIELD GROUP SUB-GROUP Ob .05 ABSTRACT (Continue on reverse if necessary at the purpose of the Air force He the aerial spraying of herbit experienced increased deaths a Baseline Mortality Report was rebruary 1984, and the first from the present report of the Air force He the serial spraying of herbit experienced increased deaths a Baseline Mortality Report was rebruary 1984, and the first from the present report of the present report of the present report of the present report of the present mortality between the present mortality between the sufficient time may have a sufficient time may have	18. SUBJECT TERMS (C) Epidemiologic I Matched Cohort Hortality Study and identify by block acts salth Study is t loides in Vieta is a result of released in Ju collow-up mortal at could be con- leaching died. Ity analyses die in the subjects were de ots had died. Ity analyses die in the subject of a	1985 Novementinue on reverse nvestigation Design; Noncomber) to determine during their particular study in clusively at ird mortality study in termined. Death contains and comparison one clinical structure of the contained of	if necessary and; Air Force oncurrent Possible Ranch dipation in a Baseline December tributed to y analyses. In of 31 December of 31 December on groups. I relatively significates with the Ranch of STORTY CLASSIFICATION SECURITY CLASSIFICATION SEC	identify by block Health Sturospective dividuals in Hand operation that programmed that programmed that programmed the peaker 1984 are obtained to the continued by young and loant continued	ok number) idy; Design; nvolved in ation have gram. The Report in ther study for dioxin in the 1257, 55 Ranch ed on all ignificant mortality ditions to

Air Force Health Study Mortality Update - 1985

EXECUTIVE SUMMARY

BACKGROUND

The purpose of the Air Force Health Study is to determine whether those individuals involved in the aerial spraying of herbicides in Vietnam during the Ranch Hand operation have experienced any adverse health effects as a result of their participation in that program. The study evaluates both mortality (death) and morbidity (disease) in these individuals over a 20-year period after the studies were initiated.

The Baseline Mortality Report was released in June 1983, the Baseline Morbidity Report in February 1984, and the first follow-up mortality study in December 1984. Neither study demonstrated health effects which could be conclusively attributed to herbicide or dioxin exposure. The reader is referred to reports of the studies for further details (1, 2, 3).

METHOD

The present report describes the third mortality analyses. Deaths in the 1257 Ranch Hand and 6171 Comparison subjects were determined, using the data sources of the Air Force, Veterans Administration, Social Security Administration, Internal Revenue Service, and personal contacts. As of 31 December 1984, 55 Ranch Handers and 285 Comparison subjects had died. Death certificates were obtained on all subjects.

Extensive statistical analyses were accomplished, as detailed in the report, to compare the death experience in the Ranch Hand population with the Comparison group. In addition, death experience in these groups was compared to the 1978 U.S. White male mortality experience, the 1978 Department of Defense Nondisability Retired Life Table, and the active U.S. civil service population as discussed in the 1984 mortality report (3). The West Point class of 1956 and the active duty USAF population are not appropriate groups for comparing to the study population and, consequently, they have not been used in the analyses in this report.

RESULTS

As was the case in the last mortality report, the current mortality analyses did not reveal any statistically significant differences in mortality between the exposed and Comparison groups. The percentages dead in each major category are summarized below. Within categories of rank and occupation none of the differences between the Ranch Hand and Comparison groups are statistically significant.

	Percen Ranch Hand	t Deaths Comparison
Rank		
Officers	3.4	4.3
Enlisted	4.9	4.8
Occupation		
Flying	3.7	5.1
Ground	5.1	4.1
	Ranch Hand	Comparison
Total	·	
Overall	4.4	4.6

As was reported in the 1984 mortality study, the Ranch Hand officers had a nonstatistically significant though slightly lower death rate than their Comparisons. There is an interaction in these data, however. Ranch Hand officers born between 1905 and 1935 have experienced fewer deaths than Comparison officers born during the same era. On the other hand, Ranch Hand officers born after 1935 have experienced more deaths than their Comparisons. Although these differences within birth-year strata are not statistically significant, this change in the group by survival status relationship with birth year is statistically significant. Additionally, Ranch Hand officers experienced fewer deaths after age 35 years than did Comparison officers, while Ranch Hand officers experienced more deaths before age 35 years than did Comparisons. The relevance of these observations is unclear at this time.

Ranch Hand flyers had a nonstatistically significant though slightly lower death rate than Comparisons, and Ranch Hand ground personnel had a slightly higher but nonstatistically significant death rate than the Comparisons.

The herbicide/dioxin exposure index described in the morbidity report was applied to the data, and no relationship between exposure and mortality experience was identified.

As was also noted in the 1984 mortality study, analyses consistently demonstrated significantly better survival in the Ranch Hand officers than Ranch Hand enlisted members, as was the case with Comparison officers and Comparison enlisted personnel. Cause-specific analyses did not demonstrate any increased Ranch Hand mortality for accidents, suicide, homicide, malignancy or circulatory system disease. No unusual patterns of malignancy were observed in either the Ranch Hand or Comparison groups, a finding which would be expected from the small number of deaths to date. When compared to the 1978 U.S. White male population, all subgroups are living longer than expected. All groups had a mortality experience similar to the civil service population.

CONCLUSION AND RECOMMENDATION

Continued mortality surveillance is recommended, since the study groups are still relatively young and healthy. While sufficient time may have elapsed for some clinically significant conditions to occur, additional time is necessary for other conditions, which may possibly be attributable to herbicide exposure, to develop. At this time, however, there is no evidence of increased mortality as a result of herbicide exposure in those individuals who accomplished the Ranch Hand spray operations in Vietnam.

Table of Contents

		Page
Exe	ecutive Summary	i
Tab	ole of Contents	iv
1.	Introduction	1
2.	Ranch Hand Versus Comparison Group Analyses	2
3.	Within-Group Analyses of Mortality	13
4.	Cause-Specific Analyses	14
5.	Noncause-Specific Comparisons with External Population	21
6.	Further Covariate Adjustments	32
7.	Statistical Aspects	34
8.	Future Commitments	37
9.	Summary and Conclusions	37
	References	39
	Principal Investigators	41
	Appendix - Mortality by Year-of-Birth	42

1. Introduction

This report updates the findings of the last mortality report (3) released on 10 December 1984. The reader is referred to the baseline mortality report (1), released on 30 June 1983, for information regarding the study design, statistical procedures and the mortality determination process. One newly identified non-Black enlisted-ground Ranch Hander has been added to the data file since the last report. This individual was previously known, but confirmation of his eligibility was delayed. Summary counts of the population at risk and the number of deaths in each of the two groups (Ranch Hand and Comparison) stratified by rank and occupation are shown in Table 1. The analyses in this report are based on this data and the data in Table 4. Table 2 contains the counts of new deaths in the population since the Table 3 in this report corresponds to Table 3 in the baseline report last report. and contains summary counts and death rates by occupation, race and group. December 1984 report, the mortality experience of the study population was contrasted with data from West Point graduates and the active duty Air Force population. As noted in that report, the West Point group consists only of current and former officers, and with respect to the active duty Air Force population, individuals with serious illness are generally not allowed to remain on active duty. contrasts with these groups are not appropriate in the context of this study and have not been used. All tables in this report correspond to similar tables in the last annual report. These counts reflect cumulative mortality as of 31 December 1984 (certified as of 15 April 1985).

Careful interpretation of the findings in this and previous reports in this series requires consideration of the large sample approximations and assumptions associated with the statistical procedures. Current knowledge regarding these statistical aspects is presented in Section 7 of this report.

Table 1
Summary Counts of Death by Rank and Occupation

•	Ranch Hand			Comparison			
Rank	At Risk	<u>Dead</u>	<u>Rate (%)</u>	At Risk	Dead	Rate (%)	
Officers Enlisted	466 791	16 39	0.034 (3.4) 0.049 (4.9)	2278 3893	98 187	0.043 (4.3) 0.048 (4.8)	
Occupation							
Flying Ground	646 611	24 31	0.037 (3.7) 0.051 (5.1)	3163 3008	161 124	0.051 (5.1) 0.041 (4.1)	
Total	1257	55	0.044 (4.4)	6171	285	0.046 (4.6)	

In Table 2, the number "at risk" is the number alive on 1 January 1984.

Table 2
Deaths During 1984 by Rank and Occupation

		Ranch Har	ıd	Comparison			
Rank	At Risk	1984 Deaths	Rate Per 100	At Risk	1984 Deaths	Rate Per 100	
Officer Enlisted	451 752	1 0	0.2	2187 3719	7 13	0.3 0.3	
Occupation							
Flying Ground	623 580	1 0	0.2	3014 2892	12 8	0.4 0.3	
Total	1203	1	0.1	5906	20	0.3	

Since so few deaths have occurred during 1984, the statistical findings and interpretations presented in this report are very similar to those in the 1984 mortality update (3).

Table 3
Occupational and Race-Specific Mortality

			Ranch Ha	and	(Compart:	sons
Race	Occupation A	t Risk	Dead	Rate Per 100	At Risk	Dead	Rate Per 100
Non-Black	Officer-Pilot	350	12	3.4	1740	79	4.5
	Officer-Nav	82	3	3.7	390	15	3.8
	Officer-Other	25	1	4.0	123	4	3.3
	Enlisted-Flt En	g 191	7	3.7	93 5	57	6.1
	Enlisted-Other	533	28	5.3	2 62 8	108	4.1
Black	Officer-Pilot	6	0	0.0	13	0	0.0
	Officer-Nav	2	0	0.0	10	0	0.0
	Officer-Other	1	0	0.0	2	0	0.0
	Enlisted-Fit En	q 15	2	13.3	75	10	13.3
	Enlisted-Other	52	_2	3.8	255	<u>12</u>	4.7
•	Total	1257	55	4.4	6171	285	4.6

2. Ranch Hand Versus Comparison Group Analyses.

Survival contrasts were made using linear rank procedures, survival curves, relative risk estimation and standardized mortality ratios. Survival curves were estimated by the product-limit estimate of Kaplan and Meier (4). Linear rank testing was carried out using the logrank test and Prentice's censored data extension of the Wilcoxon test (5). All linear rank tests were carried out with matched sets merged when Ranch Handers differed by less than one year relative to date of

birth. Within each stratum of job and race, these merged matched sets were used as separate strata for testing purposes. The matched data relative risk procedure, due to Ejigou and McHugh (6), is applied only to the 1241 Ranch Handers with matched Comparisons, and the stratified relative risk or SMR estimate is applied to all 1257 Ranch Handers.

Group contrasts were made on officers, enlisted personnel, flying personnel, ground personnel and the total group. Summary counts are shown in Table 4.

Table 4
Summary Counts by Rank, Occupation and Group

Flying Personnel

<u>Officer</u>		<u>En</u>	Enlisted			<u>Total</u>			
Groups	At Risk	Dead	Rate Per 100	At Risk	Dead	Rate Per 100	At Risk	Dead	Rate Per 100
Ranch Hand	440	15	3.4	206	9	4.4	646	24	3.7
Comparisons	2153	94	4.4	1010	67	6.6	3163	161	5.1

Ground Personnel

		Offic		<u>En</u>	listed		To	tal	
Groups	At Risk	<u>Dead</u>	Rate Per 100	At Risk	Dead	Rate Per 100	At Risk	<u>Dead</u>	Rate Per 100
Ranch Hand Comparisons	26 125	1 4	3.8 3.2	585 2883	30 120	5.1 4.2	611 3008	31 124	5.1 4.1

Survival curves were estimated only for officers, enlisted, flying, ground personnel and all personnel in Ranch Hand and Comparison groups. There is a substantial degree of overlap between these subgroups, with 96% of both the Ranch Hand and Comparison ground personnel being enlisted. The enlisted category includes both ground support and flying enlisted personnel. Survival curves for the overall Ranch Hand and Comparison groups are shown in Figure 1. The curves for officers, enlisted, flyers and ground personnel are shown in Figures 2 through 5.

Figure 1
Survival Curve Estimates for All Ranch Handers and All Comparisons

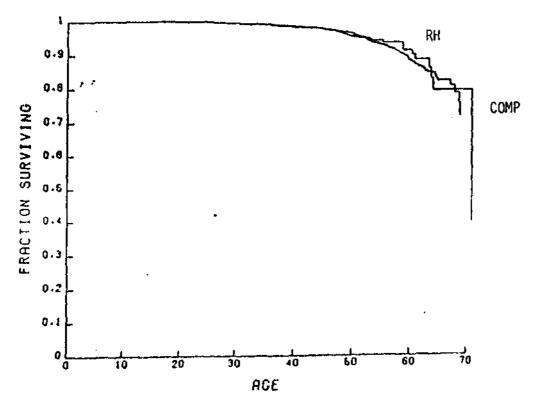


Figure 2

Survival Curve Estimates for Ranch Hand and Comparison Officers

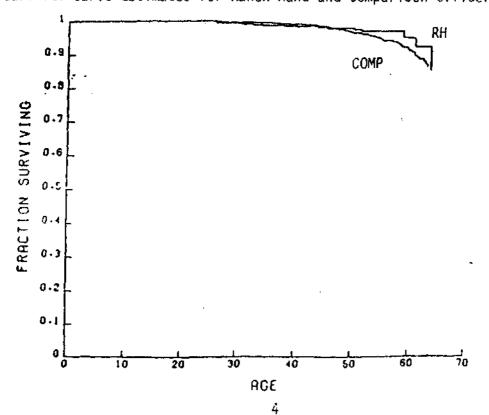


Figure 3
Survival Curve Estimates for Enlisted Ranch Handers and Comparisons

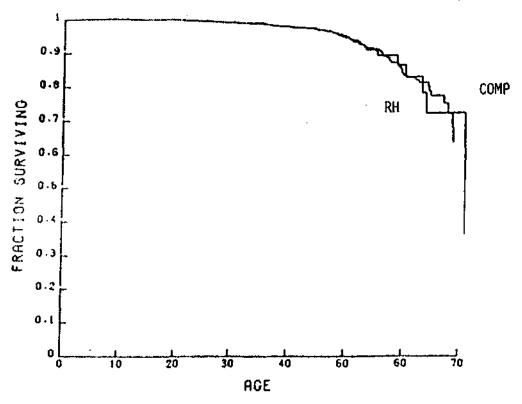


Figure 4
Survival Curve Estimates for Ranch Hand and Comparison Flyers

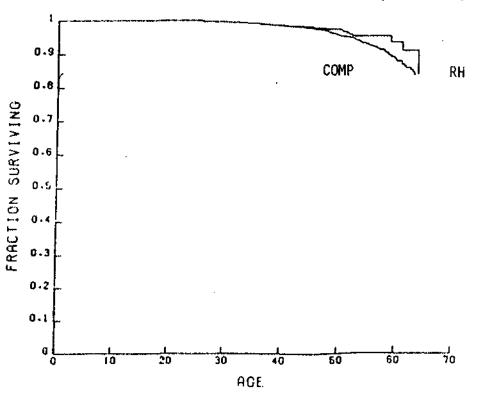
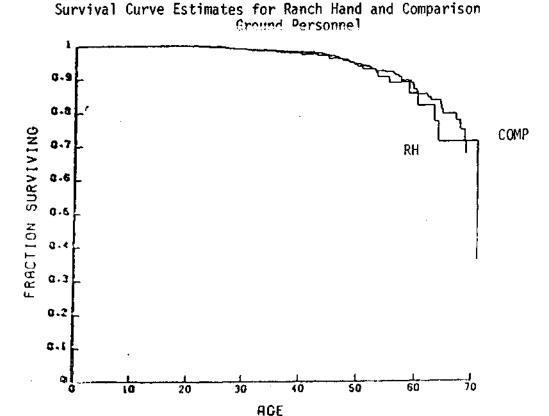


Figure 5



The patterns qualitatively evident in these graphs are seen quantitatively in subsequent statistical analyses.

Linear rank procedures were carried out on the same four subgroups and on all personnel to assess death patterns by time. These procedures are designed so that the statistic will be positive when Ranch Handers are dying before Comparison subjects and negative when Comparisons are dying prior to Ranch Handers. The results are shown in Table 5 (Table 6 in the baseline report).

The linear rank statistic used is a valid measure of group difference only when this difference occurs consistently across strata. Since the strata in these analyses were formed by date-of-birth, occupation and race, the linear rank statistic is valid only when the direction of the group difference in death times does not change with date-of-birth, race and occupation. As discussed in Section 7, there is currently no statistical procedure available for testing the assumption that differences in group survival distributions remain constant across strata. As will be shown later, there is an indication, however, that there is an effect of date-of-birth on relative risks in the officer subgroup. Thus, the logrank and Wilcoxon tests on officers must be interpreted carefully. However, these data suggest that the summary statistics for the remaining subgroups are valid. Further, since there is an indication that mortality contrasts change with rank and occupation, the overall (tatal) logrank and Wilcoxon values and p-values, shown in Table 5 are not valid summary statistics.

Table 5
Test Results and P-Values for Noncause-Specific Survival

	Logr	ank	Wilcoxon			
Group	(Value)	P-Value	(Value)	P-Value		
Officer	(-0.835)	0.40	(-0.903)	0.37		
Enlisted	(0.187)	0.85	(0.161)	0.87		
Flying	(-1.34)	0.18	(-1.42)	0.16		
Ground	(0.976)	0.33	(0.093)	0.34		
Total	(-0.305)	0.76	(-0.344)	0.73		

Table 5 suggests that ground personnel in the Ranch Hand group are dying sooner than their matched Comparisons (logrank = 0.976), but again the difference is not statistically significant (p=0.33). The negative values of the logrank and Wilcoxon statistics for officers (logrank = -0.835) and flying personnel (logrank = -1.34) suggest that Ranch Handers in this group may be living longer than their matched Comparisons, but not to a statistically significant degree.

Similar analyses on the same subgroups (officer, enlisted, flying, ground and total) were carried out on data from non-Black subjects only. The results are shown in Table 6.

Table 6

Test Results and P-Values for Noncause-Specific Survival Non-Black Ranch Handers and Non-Black Comparisons

	Logr	ank	Wilcoxon		
Group	(Value)	P-Value	(Value)	P-Value	
Officer Enlisted Flying Ground	(-0.819) (0.211) (-1.43) (1.10)	0.41 0.83 0.15 0.27	(-0.885) (0.192) (-1.50) (1.08)	0.38 0.85 0.13 0.28	
Total	(-0.286)	0.78	(-0,320)	0.75	

The findings in Table 6 clearly parallel those of Table 5, as would be expected from the small size of the Black cohort in this study.

Relative risk estimates, the associated 95% confidence intervals, two-sided p-values for testing the null hypothesis of relative risk equal to unity, and power for detecting a relative risk of 2 in these data are shown in Table 7. These estimates are based on a matched data algorithm and summarize the relative prevalence of death in the Ranch Hand and Comparison groups. The estimated relative risks are valid summary statistics only when relative risk can be assumed to be constant across date of birth strata. Again, there is indication that this assumption is not met in the officer cohort so their estimated relative risks must be viewed with caution. On the other hand, the assumption appears to be

met in the flying, ground and enlisted subgroups so these relative risk estimates do appear to be valid. Similarly, since there is an indication that relative risk changes with rank and occupation, the overall relative risk, 0.915, is not a valid summary statistic.

Table 7

Relative Risks, 95% Confidence Intervals, P-Values and Power for Noncause-Specific Deaths to Date (1241 Ranch Handers Versus 6171 Matched Comparisons)

Group	Rel Risk	Conf Int	P-Value	Power
Officer Enlisted Flying Ground	0.715 0.987 0.692 1.21	(0.311, 1.12) (0.622, 1.35) (0.377, 1.01) (0.708, 1.72)	0.26 0.94 0.12 0.35	0.90 0.99 0.98 0.94
Total	0.915	(0.636, 1.20)	0.57	1.0

Table 7 shows that Ranch Hand flyers are experiencing fewer deaths than their matched Comparisons (relative risk = 0.692), but this group difference is not statistically significant (p=0.12). The Ranch Hand ground personnel experienced more deaths (relative risk = 1.21) than their matched ground Comparisons, but again, this excess is also not statistically significant (p=0.35). The statistical power to detect a relative risk of two is quite strong (equal to or greater than 90%).

Year-of-birth specific mortality rates are given in Tables 8 through 12, with the corresponding standardized mortality ratios (SMR) and associated p-values (7). In each analysis, the Comparison group is the internal standard. The SMR will accurately estimate the relative risks within each stratum in these analyses if the year-of-birth specific relative risks are equal. A likelihood ratio test for the hypothesis of, equal year-of-birth specific relative risks was carried out for each analysis, and its p-value is denoted by P1. In addition, the hypothesis that the relative risk is unity, given that relative risk is constant across strata, was tested; its p-value is denoted by P2. The SMR and both p-values are given for each contrast. Additional analyses were conducted and are presented at the end of this section. They indicate that the hypothesis of equal year-of-birth specific relative risks may not be met in the officer cohort.

Table 8

Year-Of-Birth Specific Mortality Rates
(1257 Ranch Handers Versus 6171 Comparisons)
(SMR = 0.954, P1 = 0.22, P2 = 0.73)

	Comparison					
Birth			Rate		•	Rate
Year	At Risk	Dead	Per 100	<u>At Risk</u>	Dead	<u>Per 100</u>
1905-1914	5	2	40.0	14	3	21.4
1915-1919	17	5	29.4	96	14	14.6
1920-1924	48	3	6.3	241	30	12.4
1925-1929	84	2	2.4	501	46	9.2
1930-1934	305	18	5.9	1389	79	5.7
1935-1939	211	7	3.3	1020	39	3.8
1940-1944	210	5	2.4	1096	24	2.2
1945-1954	377	<u>13</u>	3.4	1814	50	2.8
Total	1257	55		6171	285	

Table 9

Officer-Specific Mortality Rates by Year-Of-Birth (SMR =0.791, P1 = 0.41, P2 = 0.37)

	Ra	nch Hai	nd	Comparison			
Birth	-		Rate		·	Rate	
<u>Year</u>	At Risk	Dead	<u>Per 100</u>	At Risk	<u>Dead</u>	Per 100	
1910-1924	41	3	7.3	205	21	10.2	
1925-1934	194	5	2.6	930	52	5.6	
1935-1939	95	4	4.2	458	13	2.8	
1940-1944	91	2	2.2	495	7	1.4	
1945-1949	45	2	4.4	190	_5	2.6	
Total	466	16		2278	98		

Table 10
Enlisted-Specific Mortality Rates by Year-Of-Birth (SMR = 1.03, P1 = 0.67, P2 = 0.89)

	Ranch Hand			Comparison			
Birth			Rate		•	Rate	
<u>Year</u>	At Risk	<u>Dead</u>	Per 100	At Risk	Dead	Per 100	
1905-1914	4	2	50.0	12	3	25.0	
1915-1919	9	2	22.2	54	10	18.5	
1920-1924	16	3	18.8	80	13	16.2	
1925-1929	41	2	4.9	211	26	12.3	
1930-1934	154	13	8.4	749	47	6.3	
1935-1939	116	3	2.6	562	26	4.6	
1940-1944	119	3	2.5	601	17	2.8	
1945-1954	332	11	3.3	1624	45	2.8	
Total	791	39		3893	187		

Table 11

Flying-Specific Mortality Rates by Year-Of-Birth (SMR =0.726, P1 = 0.85, P2 = 0.13)

	Ra	nch Hai	nd	Comparison		
Birth			Rate		·	Rate
<u>Year</u>	At Risk	<u>Dead</u>	Per 100	At Risk	Dead	Per 100
1915-1924	44	4	9.1	220	26	11.8
1925-1934	272	10	3.7	1316	84	6.4
1935-1939	145	6	4.1	698	26	3.7
1940-1944	121	2	1.7	653	15	2.3
1945-1949	64	_2	3.1	276	10	3.6
Total	646	24		3163	161	

Table 12

Ground-Specific Mortality Rates by Year-of-Birth (SMR = 1.23, P1 = 0.59, P2 = 0.33)

		Ranch			Compar	_
Birth <u>Year</u>	At Risk	<u>Dead</u>	Rate Per 100	At Risk	Dead	Rate Per 100
1905-1914	5	2	40.0	14	3	21.4
1915-1919	8	1	12.5	51	8	15.7
1920-1924	.13	. 3	23.1	66	10	15.2
1925-1929	31	2	6.5	151	20	13.2
1930-1934	86	8	9.3	423	21	5.0
1935-1939	66	1	1.5	322	13	4.0
1940-1944	89	3	3.4	443	9	2.0
1945-1954	313	11	3.5	<u>1538</u>	40	2.6
Total	611	31		3008	124	

Additional log-linear analyses of the data in Tables 9 through 12 were carried out. These analyses are directed at the hypothesis already tested and reported, via the p-value (P1), but have an advantage in that they are more powerful. They have a disadvantage in that, since they were carried out after the data had already been tested, the overall level of significance is higher than the nominal 5%. The extent of the increases in power and significance level is not known. When year-of-birth is dichotomized (1905-1934, 1935-1954) and survival status (alive, dead) is analyzed by group (Ranch Hand, Comparison) and rank (officer, enlisted), a borderline significant four-way interaction is evident (p=0.054). The officer and enlisted relative risks are 0.53 and 1.10 in the 1905-1934 year-of-birth stratum and 1.58 and 0.95 in the 1935-1954 birth-year stratum. There were no three-way interactions in this analysis. When rank is replaced by flying status (flying, ground) in this four-factor analysis, no four-way interaction is seen (p=0.085), and no significant group by flying status by birth-year interaction (p=0.92) is observed.

Further, when the officer, enlisted, flying and ground subgroups are analyzed separately on survival status, group and birth-year, there is no three-way interaction for enlisted (p=0.67), flying (p=0.30) or ground personnel (p=0.28) but there is a significant three-way interaction for the officers (p=0.044). That is, the survival status by group relationship changes with year-of-birth in the officer cohort. Two-factor p-values are 0.87 for enlisted, 0.12 for flying, and 0.077 for ground personnel. These findings are consistent with previous analyses.

Taken together, these log-linear analyses suggest that relative risk changes with year-of-birth in the officer cohort. Specifically, the overall death experience of the Ranch Hand officers appears to compare favorably with the Comparisons. However, these diminished death rates appear to be found in the Ranch Hand officers born before 1935, while Ranch Hand officers with later birth dates evidence a rate equal to or exceeding that of the Comparisons (as seen in Table 14).

These findings cast doubt upon the validity of the SMR and, possibly, the linear rank procedures, as summary statistics for the officer cohort. The SMR appears to be a valid summary statistic for Ranch Hand and Comparison contrasts within the enlisted, flying and ground cohorts.

A summary of logrank, relative risk and SMR results obtained is shown in Table 13.

Table 13
Noncause-Specific Statistical Summary

Age at Death

	Logrank			
Group	Value	P-Value		
Officer Enlisted Flying Ground	-0.835 0.187 -1.34 0.976	0.40 0.85 0.18 0.33		
Total	0.305	0.76		

Deaths to Date

Group	RR	P-Value	SMR	P-Value
Officer Enlisted Flying Ground	0.715 0.987 0.692 1.21	0.26 0.94 0.12 0.35	0.791 1.03 0.726 1.23	0.37 0.89 0.13 0.33
Total	0.915	0.57	0.954	0.73

The data in Table 13 show reasonable consistency. The ground cohort displays excess death in the Ranch Hand group in contrast to the Comparison group, and the Ranch Hand flying cohort exhibits fewer deaths, but again these group differences are not statistically significant. The officer cohort evidences less death in the Ranch Hand group in contrast to the Comparison group but, again, this group difference is not statistically significant. However, as discussed above and shown in Table 14, these data appear to suggest that favorable mortality experience occurs in those officers born before 1935, while Ranch Hand officers born after 1935 appear to have experienced the same or greater death rate than their Comparisons.

Table 14

Death Rates by Group, Rank, Occupation and Year-of-Birth

Rank	Year-of-Birth	Death Rate Ranch Hand	per 100 Comparison	Relative Risk
Officer	1905-1934	3.4	6.4	0.53
	1935-1954	3.5	2.2	1.58
Enlisted	1905-1934	9.8	9.0	1.10
	1935-1954	3.0	3.2	0.95
Occupational	Year-of-Birth	Death Rat Ranch Hand	te per 100 Comparison	Relative Risk
Flying	1905-1934	4.4	7.2	0.62
	1935-1954	3.0	3.1	0.97
Ground	1905-1934	11.2	8.8	1.27
	1935-1954	3.2	2.7	1.19

The favorable, though not statistically significant, survival experience of Ranch Hand flying personnel, relative to the matched Comparison flyers is shown in Figure 4, where the survival curves for Ranch Hand and Comparison flyers are drawn on the same scale and coordinate system. In contrast, the relatively poorer, but not statistically significant, survival experience of the Ranch Hand ground personnel is illustrated in Figure 5, wherein the Ranch Hand and Comparison ground personnel survival curves are drawn on the same coordinate system.

3. Within-Group Analyses of Mortality

Within-group year-of-birth adjusted contrasts by occupation and rank via SMR's are summarized in Table 15. The data supporting these SMR analyses are shown in Appendix Tables 1 through 4.

Table 15
Summary of Within-Group SMR Analyses

Subgroups	SMR	<u>P1</u>	<u>P2</u>
Officers versus Enlisted Ranch Hand	0.515	0.27	0.047
Comparison	0.648	0.88	0.001
Flying versus Ground			
Ranch Hand Comparison	0.572 0.909	0.41 0.46	0.067 0.65

Table 15 shows that Ranch Hand officers are having significantly fewer deaths (SMR=0.515, p=0.047) than Ranch Hand enlisted personnel, after adjustment for year-of-birth. This officer versus enlisted differential is also significant and in the same direction in the Comparison group (SMR=0.648, p=0.001). The table also suggests a favorable mortality experience of Ranch Hand flyers relative to that of the Ranch Hand ground personnel (SMR=0.572, p=0.067), although this difference is not statistically significant. A flyer versus ground differential is not apparent in the Comparison group (SMR=0.909, p=0.65).

4. Cause-Specific Analyses

Table 16 shows death counts by cause and subgroup (flying officer, ground officer, flying enlisted and ground enlisted). Counts are shown for all 1257 Ranch Handers and the 6171 Comparisons. The distribution of new deaths in the Ranch Hand and Comparison groups are presented in Table 17, and age-adjusted relative risks for these data are shown in Table 18. Relative risks are calculated using a matched data algorithm; hence, only the 1241 Ranch Handers having matched Comparisons are used. Of the 16 unmatched Ranch Handers, two have died; a flying officer died of an accident and a ground airman died of circulatory system disease. Since these data are sparse, relative risks are only calculated on officer, enlisted, flying and ground subgroups, as well as on the total population.

One cell in Table 18, the analysis of malignancy deaths in flying personnel, contains two p-values for the significance of the relative risk estimate. The first is calculated using a null variance of the estimated relative risk and the second, within parentheses, is calculated using the non-null variance estimate. A null variance is defined as a variance derived upon the assumption that the true relative risk is unity. A non-null variance is derived without any assumption about the true value of the relative risk. The choice of which variance estimate to use in the standardization of the test statistic is currently a point of research in theoretical statistics. We have chosen to use the null variance when computing p-value because of analogies with other testing situations and because our power studies have shown the resulting test to be more powerful than the test using the general non-null estimate. Unfortunately, the non-null variance must be used in computing 95% confidence intervals for the relative risk, making the p-value and confidence interval sometimes incompatible.

Table 16

Deaths by Cause and Subgroup

		0f	ficer			En1	isted			
	Fly	ing	Gro	und	FTy	ing	Gro	und	<u>Tot</u>	<u>a1</u>
Cause	RH	<u>c</u>	RH	<u>c</u>	RH	$\overline{\mathbf{c}}$	RH	$\overline{\mathbf{c}}$	RH	<u>c</u>
Accident	8	33	0	1	4	27	7	35	19	96
Suicide	0	5	1	1	1	3	1	9	3	18
Homicide	0	0	0	0	0	2	2	3	2	5
Parasitic infection	0	2.	0	0	0	0	0	2	0	4
Malignant neoplasm	0	15	0	1	1	14	5	21	6	51
Uncertain neoplasm	0	1	0	0	0	.0	0	1	0	2
Endocrine	0	1	0	0	0	0	1	0	1	1
Mental disorder	0	0	0	0	0	0	0	1	0	1
Nervous System	0	1	0	0	0	0	0	1	.0	?
Circulatory System	5	28	0	0	1	14	12	38	18	80
Respiratory System	0	2	0	0	0	3	0	2	0	7
Digestive System	2	4	0	1	1	3	2	5	5	13
Genitourinary System	0	1	0	0	0	0	0	2	0	3
Congenital anomalies	0	0	0	0	0	0	0	0	0	0
Ill defined	_0	_1	_0	_0	_1	_1	_0	_0	_1	2
Total	15	94	1	4	9	67	30	120	55	285

Table 17
New Deaths by Cause

Cause	Ranch Hand	Comparison
Accident	0	2
Suicide	0	2
Homicide	0	1
Malignant Neoplasm	0	. 8
Circulatory System	1	5
Respiratory System	0	2
Totals	1	20

Table 18

Cause-Specific Age Adjusted Relative Risks by Group (1241 Ranch Handers versus 6171 Comparisons)

Group	<u>Ştatistic</u>	Accident	Suicide	<u>Cause</u> Homicide	Malignancy	<u>Circulatory</u>	Digestive
Officer	Rel Risk Conf Int P-Value Power	0.968 (0.160, 1.78) 0.94 0.61	0.833 (0.000, 2.60) 0.87 0.29			0.577 (0.000, 1.27) 0.38 0.54	2,00 (0,000, 5,280) 0,36 0,28
Enlisted	Rel Risk	0.830	0.833	2.00	0.834	1.08	1.88
	Conf Int	(0.275, 1.38)	(0.000, 2.08)	(0.000, 5.28)	(0.104, 1.57)	(0.361, 1.78)	(0.000, 4.36)
	P-Yalue	0.59	0.81	0.36	0.69	0.82	0.31
	Power	0.79	0.37	0.28	0.62	0.72	0.32
Flying	Rel Risk Conf Int P-Value Power	0.905 (0.316, 1.49) 0.77 0.78	0.625 (0.000, 1.92) 0.67 0.32		0.172 (0.000, 0.516 0.069(0.000) 0.57	0.500 6) (0.000, 1.01) 0.20 0.66	2.14 (0.000, 5.04) 0.2? 0.3!
Ground	Rel Risk	0.803	1.00	3.333	1.08	1.34	1.67
	Conf Int	(0.090, 1.52)	(0.000, 2.52)	(0.000, 9.30)	(0.009, 2.15)	(0.367, 7.31)	(0.060, 4.33)
	P-Value	0.63	1.0	0.099	0.88	0.42	0.51
	Power	0.62	0.35	0.25	0.49	0.62	0.29
Total	Rel Risk	0.917	0.833	2.00	0.579	1.02	1,92
	Conf Int	(0.447, 1.39)	(0.000, 1.85)	(0.000, 5.28)	(0.087, 1.07	(0.459, 1.58)	(0,000, 3,91)
	P-Value	0.74	0.77	0.36	0.22	0.95	0,17
	Power	0.90	0.45	0.28	0.73	0.85	0,39

^{*}P-value compatible with the confidence interval.

Tables 16 and 18 must be interpreted with care since the data are very sparse in some categories. The behavior of the Ejigou-McHugh estimate, like that of other relative risk estimates, has not been investigated when the death probabilities are very small, as is the case for the causes analyzed in Table 18. The analyses of malignant neoplasm and circulatory system deaths are more reliable than the other cause-specific analyses because these two categories contain more deaths than the others.

Digestive system mortality by ICD code is shown in Table 19, site-specific malignant neoplasm mortality is shown in Table 20 and the morphology of neoplasms is shown in Table 21. There was one case of soft tissue sarcoma in a Comparison individual, but none in the Ranch Hand group. There have been no cancer deaths in the Ranch Hand group and eight in the Comparison group during 1984.

Table 19
Digestive System Mortality

	Dear	
ICD Code	Ranch Hand	Comparison
Pancreatitis (5770)	1	2
Alcoholic cirrhosis (5712)	0	4
Nonalcoholic cirrhosis (5715)	3	3
Nonalcoholic fatty liver (5718)	0	1
Chronic liver disease (5728)	0	2
Alcoholic liver disease (5711)	1	0
Duodenal ulcer (5325)	0	1
Peptic ulcer (5334)	0	0
Hepatocellular disease (573a)	_0	_0
Total	5	13

These codes were based on death certificate data; more detailed etiologic information has been requested but not yet received for the nonalcoholic cirrhosis and fatty liver deaths. It is of interest that during 1984, there were no new deaths attributable to the digestive system in either group.

Table 20
Site-Specific Malignant Neoplasm Mortality

Site ICD Code	Ranch Hand	Comparison
Lip, oral cavity, pharynx (140-149)	0	4
Digestive organs, peritoneum (150-159)	1	12
Respiratory, intrathoracic (160-165)	3	21
Bone, connective tissue, skin,		
breast (170-175)	0	1
Genitourinary organs (179-189)	ï	3
Brain (191-192)	Õ	3
Other and ill-defined sites (195)	Ó	1
Lymphatic and hematopoietic		
tissue (200-208)	0	5
No site specification (199)	_1	_1
Total	6	51

Table 21

Morphology of Neoplasms

ICD Code 9th Ed.	Nomenclature	Deat Ranch Hand	hs <u>Comp</u>
M800	Neoplasms not otherwise specified (NOS) Brain Bronchus and Lung	0 1	1 6
	Colon	0	6 2 2 1
	Pancreas Intestinal Tract	0 .	2
	Head and Neck	0	1
M801-804	Epithelial neoplasms (NOS)		_
	Bronchus and Lung	1	10
	Esophagus	0	1
	Kidney	1 0	1
	"Nasopharynx Pancreas	0	1 1 2 0 1
	Stomach	i	õ
	Unspecified site	î	ĭ
M805-808	Papillary and Squamous Cell	_	
	Nasal Sinus	0	1
	Lip	0	1 1 1
	Tongue	0	I 1
	Lung Tonsil	. 0	1
M814-838	Adenomas and Adenocarcinomas	. 0	1
11014-030	Appendix	0	1
	Bronchus and Lung	ŏ	1 2 1 2 1
	Colon	ő	ī
	Kidney	0	2
	Stomach	0	1
	Rectum	0	1
M872-879	Nevi and Melanomas		
	Skin (NOS)	0	1
моос	Mediastinal	1	0
M905	Mesothelioma Bronchus and Lung	0	1
M938-948	Gliomas	C)	J
11200 240	Frontal Lobe	0	1
•	Brain (NOS)	0	ï
M959-963	Lymphomas NOS and Diffuse		
	Lymphomas (NOS)	0	1
M964	Reticulosarcoma	0	1
MUCE UCC	Malignant lymphoma histiocytic, (NOS)	0	1
M965-966	Hodgkin's disease Hodgkin's (NOS)	O	2
M986	Myeloid Leukemias	V	« -
	Acute Myelocytic Leukemia	0	1
	y y		
	Total	6	51

5. Noncause-Specific Comparisons with External Populations

It is important to know not only how Ranch Handers and their matched Comparisons relate to each other, but also how their mortality rates compare with other military and civilian populations in the United States. These contrasts are used in an attempt to place the study groups in perspective with the overall mortality experience of known populations. Given the selection factors involved for entry to and retention in the military service, it is anticipated that the study groups would exhibit lower mortality than the U.S. White male population. Similarly, they might be expected to be more equivalent to the DOD retired personnel or occupational cohorts such as the U.S. civil service. In this report, the mortality experience of Ranch Handers and their matched Comparisons is compared with the expected death rates with reference to the 1978 U.S. White Male Life Table (8), the 1978 Department of Defense period life tables for nondisability retired military officer and enlisted personnel (9), and a 1974 U.S. active male civil service life table (10). All analyses in this section depend on the assumption, that relative risk is constant across age strata (Section 7). This assumption is not currently testable.

5.1 Comparisons with 1978 DOD Life Tables

In Tables 22 and 23, Ranch Hand officers and Comparison group officers are contrasted to a 1978 DOD nondisability retired officer life table and in Tables 24 and 25, Ranch Hand and Comparison group enlisted personnel are compared with a 1978 DOD nondisability retired enlisted life table. In each table, the column labeled "At Risk" lists the number of subjects entering each five-year age interval, the column labeled "Deaths" tabulates the number of deaths in the age intervals and the column labeled "Expected Deaths" gives the expected number of deaths in the age intervals of the study subjects if they had experienced the same death rates as those specified by the DOD table. The value of the test statistic for testing the null hypothesis of equality of the study and referenced life table is denoted by T; its two-sided p-value is denoted by P. While each table summarizes the findings with five-year age intervals for ease of presentation, one-year age intervals were used for the actual computation of the statistic T. A negative value of T means that the study cohort has lived longer than expected relative to the reference population. The magnitude of the statistic T is sample-size dependent. All contrasts are unadjusted for race since the DOD tables are not race-specific. analyses are conditioned on survival to age 35, since active duty personnel are not eligible for retirement prior to that age and, therefore, the DOD tables do not begin until that age. The totals in Tables 22 through 25 do not, therefore, agree with Table 1.

Table 22

Ranch Hand Officer Versus DOD Nondisability Retired Officer Life Table (T=-4.43, $P \le 0.001$)

<u>Age</u>	At Risk	Deaths	Expected Deaths
37-39	459	2	2.8
40-44	414	1	4.7
45-49	324	1	5.4
50-54	232	2	4.7
55-59	84	1	2.7
60-64	40	2	1.7
65-69	6	0	0.2
70-70	1	0	0.0
Total		9	22.2

Table 23 Comparison Officers Versus DOD Nondisability Retired Officer Life Table (T=-3.71, P < 0.001)

<u>Age</u>	At Risk	<u>Deaths</u>	Expected Deaths
35-39	2264	12	22.2
40~44	2067	14	23.1
45-49	1565	25	25.5
50-54	1095	15	23.0
55-59	472	10	13.9
60-64	192	8	8.2
65-69	40	0	1.9
70-70	2	0	0.0
Total		84	117.9

Table 24

Ranch Hand Enlisted Personnel Versus DOD Nondisability
Retired Enlisted Life Table
(T=-1.01, P=0.31)

<u>Age</u>	At Risk	Deaths	Expected Deaths
35~39	771	7	8.4
40-44	454	5	6.4
45-49	333	6	7.9
50-54	214	6	6.5
55-59	67	2	3.0
60-64	26	3	2.0
65-69	10	0	1.0
70-71	3	_1	0.2
Total		30	35.5

Table 25

Comparison Enlisted Personnel Versus DOD Nondisability Retired Enlisted Life Table (T=-4.29, P<0.001)

Age	At Risk	Deaths	Expected Deaths
35-39	3777	21	41.2
40-44 45-49	2233 1 6 28	21 38	31.3 38.7
50-54 55-59	1054 331	21 16	32.1 14.8
60-64	130	6	10.5
65-69 70-74	57 9	4 0	5.0 1.2
75-76	2	0	0.2
Total		127	175.1

Tables 22 and 23 show highly favorable mortality experiences for Ranch Hand and Comparison officers. Conditioned on survival to age 35, they are living significantly longer than expected using the DOD death rates (p 0.001 for both groups). Tables 24 and 25 show that Ranch Hand enlisted personnel are experiencing mortality patterns similar to the DOD retired enlisted population (p=0.31), and the Comparison enlisted personnel are living significantly longer (p 0.001) than the DOD nondisability retired enlisted population. This, together with the nonsignificant logrank value for Ranch Hand versus Comparison enlisted personnel shown in Table 5 (p=0.85), suggests that the Ranch Hand versus Comparison contrasts may change with age at death. A view of this is seen in Table 26, which shows linear rank test results, comparing Ranch Handers and Comparisons conditioned on survival to age 35 (analogous to Table 5). Comparing the conditional analyses in Table 26 with the unconditional analyses in Table 5, it appears that group contrasts change with age at death within the officer cohort.

Table 26

Ranch Hand Versus Comparison
Test Results and P-Values for Noncause-Specific Survival
Conditioned on Survival to Age 35

Group	<u>Logrank</u>		Wilco	oxon
	(Value) P-Value		(Value)	P-Value
Officer	(-1.87)	0.061	(-1.99)	0.047
Enlisted	(0.802)	0.42	(0.810)	0.42
Flying	(-1.55)	0.12	(-1.66)	0.097
Ground	(1.12)	0.27	(1.12)	0.26
Total	(-0.481)	0.63	(-0,529)	0,60

Additional categorical analyses described below reveal the interaction suggested by the previous conditioned analyses. These are shown in Tables 27 and 28 where survival status (alive, dead) is analyzed as a function of group (Ranch Hand, Comparison) and rank (officer, enlisted) on deaths under 35-years of age and separately on deaths over 35-years of age.

Table 27

Death Before Age 35, Ranch Hand Versus Comparisons (Group By Rank By Status Interaction: P=0.044)

			Status		
Rank	Group	Alive	<u>Dead</u>	Total	Relative Risk
Officer	Ranch Hand	459	7	466	2.44
	Comparison	2264	14	2278	Z• 44
	Totals	2723	21	2744	
Enlisted	Ranch Hand	782	9	791	0.738
	Comparison	3833	<u>60</u>	3893	0.730
	Totals	4615	69	4684	

Table 28

Death After Age 35, Ranch Hand Versus Comparisons (Group By Rank By Status Interaction: P=0.039)

Status

	<u>Juan</u>						
Rank	Group	Alive	Dead	Total	Relative Risk		
Officer	Ranch Hand	450	9	459	0 520		
	Comparison	2180	84	2264	0.528		
	Totals	2630	93	2723			
Enlisted	Ranch Hand	752		1 16			
	Comparison	3706	<u>127</u>	3833	1.16		
	Totals	4458	157	4615			

In Table 27 and 28, the Ranch Hand versus Comparison contrast in the officer category is significantly different from the corresponding contrast in the enlisted category. This suggests that, among those surviving to age 35, Ranch Hand officers are experiencing fewer deaths (relative risk = 0.528) than their matched Comparisons while the Ranch Hand enlisted are experiencing more deaths than their matched Comparisons (relative risk = 1.16). This situation is reversed in those men dying prior to age 35. The relevance of these observations is unclear at this time. These death rates are summarized in Table 29. The rate that is most apparently different is the low Ranch Hand officer death rate for those officers who survived to age 35. This low rate may parallel the favorable mortality experienced by those Ranch Hand officers born before 1935, as will be shown later in this report. Further analyses in future reports will attempt to clarify these patterns.

Table 29
Death Rates by Age at Death, Group, and Rank

Death Rates per 100

Age at	Ranch Hand		Comparison	
Death	Officers	Enlisted	<u>Officers</u>	Enlisted
Before Age 35 After Age 35		1.1 (N=791) 3.8 (N=782)	0.6 (N=2278) 3.7 (N=2264)	1.5 (N=3893) 3.3 (N=3833)

5.2 Comparisons with the U.S. Active Male Civil Service Life Table

To further place the Ranch Handers and their matched Comparisons in perspective, Ranch Handers, Comparisons, and officer and enlisted personnel are contrasted with the 1974 male active U.S. civil service life table (10). These contrasts are shown in Tables 30 through 35. There was no adjustment for civil service grade in these analyses. Therefore, socioeconomic factors may not be fully equivalent, especially in the analyses of the officer and enlisted subgroups. In future mortality updates, attempts will be made to account for the grade structure of the civil service population.

Table 30

All Ranch Handers Versus U.S. Male Civil Service (T=-0.313, P=0.75)

Age.	At Risk	Deaths	Expected Deaths
21-24	1257	2	6.8
25-29	1255	7	6.0
30-34	1248	7	5.7
35-39	1230	9	7.0
40-44	868	6	8.3
45-49	657	7	9.6
50-54	446	8	7.3
55~59	151	3	3.7
60-64	66	5	2.2
65-69	16	Ō	0.6
70-71	4	_1	0.1
Total		55	57.3

Table 31
Comparison Versus U.S. Male Civil Service (T=-1.04, P=0.30)

Age	At Risk	Deaths	Expected Deaths
19-19	6171	2	10.5
20-24	6169	18	43.1
25-29	6151	29	29.5
30-34	6122	25	28.1
35-39	6041	33	34.3
40-44	4300	35	40.6
45-49	3193	63	46.4
50-54	2149	36	35.7
55-59	803	26	18.8
60-64	322	14	11.1
65-69	87	4	3.7
70-74	11	0	0.6
Total		285	302.5

Table 32

Ranch Hand Officers Versus U.S. Male Civil Service (T=-1.92, P=0.054)

Age	At Risk	Deaths	Expected Deaths
25-29	466	3	2.2
30-34	463	4	2.1
35-39	459	2	3.0
40-44	414	1	4.0
45-49	324	1	4.8
50-54	232	2	4.0
55-59	84	1	2.2
60-64	40	2	1.2
65-69	6	0	0.2
70-70	1	0	0.0
Total		16	23.7

Table 33
Comparison Officers Versus U.S. Male Civil Service (Comparisons: T=-1.88, P=0.060)

Age	At Risk	Deaths	Expected Deaths
25-29	2278	9	10.9
30-34	2269	5	10.4
35-39	2264	12	14.6
40-44	2067	14	19.8
45-49	1565	25	22.7
50-54	1095	15	19.3
55-59	472	10	11.3
60-64	192	8	6.1
65-69	40	0	1.4
70-70	2	_0	0.0
Total		98	116.6

Table 34

Ranch Hand Enlisted Personnel Versus U.S. Male Civil Service (T=1.28, P=0.20)

Age	At Risk	Deaths	Expected Deaths
21-24	791	2.	4.3
25-29	789	4	3.8
30-34	785	3	3.6
35-39	771	7	4.0
40-44	454	5	4.3
45-49	333	6	4.8
50-54	214	6	3.3
55-59	67		1.5
60-64	26	2 3	1.0
65-69	10	0	0.5
70-71	3	1	0.1
Total		39	31.1

Table 35

Comparison Enlisted Personnel Versus U.S. Male Civil Service (T=1.54, P=0.12)

<u>Age</u>	At Risk	<u>Deaths</u>	Expected Deaths
19-19	3893	2	6.6
20-24	3891	18	27.2
25-29	3873	20	18.5
30-34	3853	20	17.6
35-39	3777	21	. 19.7
40-44	2233	21	20.9
45-49	1628	38	23.7
50-54	1054	21	16.4
55-59	331	16	7.5
60-64	130	6	5.0
65-69	57	4	2.3
70-74	9	0	0.6
Total		187	166.1

The Ranch Handers and their matched Comparisons are statistically quite close to the male civil service population. In these contrasts, the healthy worker effect is roughly equivalent although there is no adjustment for socioeconomic status. The contrasts of officer personnel in the Ranch Hand and Comparison cohorts with the male civil service reveal that the Ranch Hand and Comparison officers are experiencing a slightly, but not significantly better mortality than the civil service. Ranch Hand and Comparison enlisted personnel are experiencing more mortality than the civil service, but these differences are not statistically significant. All of these findings are consistent with the linear rank testing shown in Table 5, the relative risks in Table 6 and the SMR's in Tables 8, 9, and 10.

5.3 Comparisons with the U.S. 1978 White Male Life Table

Finally, the mortality experience of the non-Black Ranch Handers and their matched Comparisons is contrasted with the 1978 U.S. White Male Life Table.

Table 36

Non-Black Ranch Handers Versus the 1978 U.S. White Male Life Table (T=-5.63, P<0.001)

Age	At Risk	Deaths	Expected Deaths
21-24	1181	2	9.1
25-29	1179	6	9.9
30-34	1173	7	9.6
35-39	1155	8	10.7
40-44	824	5	11.6
45-49	627	7	14.5
50-54	432	7	12.4
55-59	150	3	6.7
60-64	66	5	4.7
6569	16	0	1.3
70-71	4	_1	0.2
Total		51	90.8

Table 37 Non-Black Comparisons Versus the 1978 U.S. White Male Life Table (T=-12.8, P<0.001)

<u>Age</u>	At Risk	Deaths	Expected Deaths
19-19	5816	1	10.3
20-24	5815	16	55.5
25-29	5799	27	48.6
30-34	5772	23	47.6
35-39	5693	31	53.1
40-44	4095	31	57.3
45-49	3047	56	70.1
50~54	2069	36	60.7
55~59	793	24	34.0
60~64	322	14	23.5
65-69	97	4 0	7.7
70-74	11		1.2
75-76	2	0	0.2
Total		263	469.7

Table 38 Non-Black Ranch Hand Officers Versus the 1978 U.S. White Male Life Table (T=-5.89, P<0.001)

<u>Age</u>	At Risk	<u>Dead</u>	Expected Deaths
25-29	457	3	3.8
30-34	454	4	3.7
35-39	450	2	4.7
40-44	407	ì	5.8
45-49	321	1	7.5
50-54	231	2	6.9
55-59	84	1	3.9
60~64	40	2	2.6
65-69	6	0	0.4
70-70	1	_0	0.0
Total		16	39.5

Table 39 Non-Black Comparison Officers Versus the 1978 U.S. White Male Life Table (T=-9.85, P<0.001)

<u>Age</u>	At Risk	Dead	Expected Deaths
25-29	2253	9	18.9
30-34	2244	5	18.5
35-39	2239	12	23.6
40-44	2042	14	28.8
45-49	1548	25	35.5
50-54	1086	15	33.6
55-59	472	10	20.6
60-64	192	8	12.9
6 5-69	40	0	2.8
70-70	2	0	0.0
Total		98	195.3

Table 40

Non-Black Ranch Hand Enlisted Personnel Versus the 1978 U.S. White Male Life Table (T=-2.20, P=0.028)

Age	At Risk	<u>Dead</u>	Expected Deaths
21-24	724	2	5.6
25-29	722	3	6.0
30-34	719	3	5.9
35-39	705	6	6.0
40-44	417	4	5.8
45-49	306	6	7.0
50-54	201	5	5.5
55-59	66	2	2.8
60-64	26	3	2.1
65-69	10	0	1.0
70-71	3	_1	0.2
Total		35	47.9

Table 41 Non-Black Comparison Enlisted Personnel Versus the 1978 U.S. White Male Life Table (T=-6.56, P<0.001)

<u>Age</u>	At Risk	Dead	Expected Deaths
19-19	3563	1	6.3
20-24	3562	16	34.0
25-29	3546	18	29.7
30-34	3528	18	29.0
35-39	3454	19	29.5
40-44	2053	17	28.5
45-49	1499	31	34.7
50-54	983	21	27.1
55-59	321	14	13.4
60-64	130	6	10.6
65-69	57	4	4.9
70-74	9	0	1.1
75-76	2	0	0.2
Total		165	248.9

The healthy worker effect is an expected phenomenon in these data since Air Force veterans have been selected for active duty on the basis of health and technical ability. This effect is clearly evident in the contrasts shown in Tables 36 through 41. Both Ranch Handers and Comparisons are seen to be living far longer than expected relative to the general U.S. White male population. The same effect

is seen in both Ranch Hand and Comparison officers (Tables 38 and 39) and in Ranch Hand and Comparison enlisted personnel. In contrast with previous mortality analyses (1,3), the analysis of the Ranch Hand enlisted cohort has reached statistical significance with the passage of time.

6. Further Covariate Adjustments

Some of the contrasts shown in previous sections in this report were further analyzed using information about the Vietnam experience for Ranch Handers and Comparisons. These analyses are motivated by the need for clarification of previous contrasts and should be viewed as preliminary to more complete analyses which will be presented in future reports. The information used here consists of (1) tour length and (2) a measure of cumulative exposure to dioxin. Tour length is defined as the cumulative time, in months, spent on assignment to Ranch Hand units by a Ranch Hander and to C-130 cargo units in SEA by a Comparison. Cumulative exposure to dioxin, termed the "exposure index," is defined in the Baseline Morbidity Report (2) and is proportional to the dioxin content of the herbicides being sprayed and inversely proportional to the number of persons sharing the workload with the subject to whom it is applied.

6.1 Ranch Hand and Comparison Contrasts on Tour Length

In this report, some descriptive statistics on tour length are presented. Table 42 shows the 5th, 50th, and 95th percentiles of tour length in months for flying and ground personnel, and officers and enlisted personnel in the Ranch Handers and Comparison groups. The effect of tour length on mortality will be more thoroughly investigated in future reports.

Table 42

Tour Length Percentiles (in Months) for Ranch Handers and Comparisons*

Group	<u>Rank</u>	Flying Status	5%	Percentil 50%	es <u>95%</u>	Sample <u>Size</u>	Population Size
Ranch Hand	Officer	Flying Ground	4 5	12 13	19 15	439 26	4 4 0 26
<u>Totals</u>	Enlisted	Flying Ground	4 5	12 13	21 20	206 585 1256	206 585 1257
Comparison	Officer	Flying Ground	11 11	19 18	46 43	2123 123	2153 125
	Enlisted	Flying Ground	10 10	19 18	49 45	995 2859	1010 2883
Totals						6100	6171

^{*}The totals show that one Ranch Hander and 71 Comparisons have no tour data at this time.

In general, the Comparisons had longer tour lengths than did the Ranch Handers. This is the result of longer tours of duty at noncombat zone bases (Comparisons) relative to combat area bases (Ranch Hand).

6.2 Ranch Hand Exposure Analyses

The effect of exposure on mortality was assessed on the 1140 Ranch Handers having exposure information in a log-linear analysis based on survival (dead, alive), rank (officer, enlisted), year-of-birth (1905-1934, 1935-1954), and exposure (light, medium, heavy). These data are shown in Table 43.

Table 43
Ranch Hand Mortality Adjusted for Year-Of-Birth, Rank and Exposure*

		Birth		Survi	val Status	
Exposure	Rank	Year	Dead	Alive	Total	Death Rate
						per 100
1.ight	Officer	1905-1934	1	54	55	1.8
7		1935-1954	2	61	63	3.2
	Enlisted	1905-1934	7	50	57	12.2
		1935~1954	3	121	124	2.4
Medium	Officer	1905-1934	2	79	81	2.5
		1935-1954	2	66	68	2.9
	Enlisted	1905-1934	4	51	55	7.3
		1935-1954	6	214	220	2.7
Heavy	Officer	1905-1934	5	84	89	5.6
		1935-1954	3	73	76	3.9
	Enlisted	1905-1934	6	84	90	6.7
		1935-1954		155	162	4.3
Totals			48	1092	1140	4.2

^{*117} Ranch Hand personnel either had a tour AFSC which removed any chance of exposure or were assigned to a Ranch Hand unit at a time when no spraying occurred or both. Tour information is not available for one Ranch Hander.

There is no four-way interaction (exposure/rank/birth year/survival status) in the data shown in Table 43 (p=0.40); there are no statistically significant three-way interactions involving survival, and the two-way survival by exposure interaction is not significant (p=0.54). These patterns do not indicate a herbicide exposure effect.

7. Statistical Aspects

The purpose of this section is to update the information contained in Chapter VI, Statistical Aspects, of the Baseline Mortality Report (1), regarding the properties of the statistical procedures used in this and all preceding mortality reports in this series. The procedures discussed here are: linear rank tests (5),

log-linear analysis (11), the SMR analysis (7), and the Gail and Ware study group versus reference life table analysis (12) and the Ejigou-McHugh relative risk estimator (6).

7.1 Large Sample P-Value Approximations

P-value calculations for all five of these procedures rely on large sample approximations of the distribution of the statistic under the associated null hypotheses, termed the null distribution. This is because the finite sample null distributions of these procedures have not been formulated. The relevant issue, therefore, is the adequacy of these approximations in the context of this study.

Linear rank tests: The adequacy of the large sample p-value approximation in certain linear rank procedures has been investigated via Monte Carlo simulation by Latta (13) in the two-sample situation and by Michalek, Mihalko and White (14) on one-to-many matched data. The primary goal of both of these studies was to investigate the power of certain linear rank procedures under various failure time distributions, censoring percentages and sample size configurations. In the two-sample case, the Prentice efficient score censored data extension of the Wilcoxon procedure was judged to be best overall, and in the matched data case, the logrank test with the hypergeometric variance was deemed the best overall procedure. These are the two procedures used in this and all previous mortality reports in this series. These simulation studies did not, however, attempt to assess the adequacy of the large sample distributions of these procedures as a function of sample size and percent censoring. In particular, neither study assessed the properties of these procedures with heavy censoring (as seen in these mortality data). Unpublished Monte Carlo studies conducted at the USAF School of Aerospace Medicine have shown, however, that the logrank and Wilcoxon tests achieve nominal 1% and 5% significance levels in two-sided testing on simulated 1:5 matched data with 1200 matched sets and 96% censoring when the survival distributions follow the accelerated failure time model (5) and the censoring variable is uniformly distributed. These results, while encouraging, are not directly applicable to this study since all linear rank testing in these reports were carried out with the data stratified by one-year birth intervals, race and occupation. Other simulations did confirm the validity of the large sample null distributions in this highly stratified case, but not with censoring percentages as high as 96%. Based on these published and unpublished investigations and the smallest sample sizes in this study (466 Ranch Hand officers contrasted with 2278 Comparison officers), the authors of this report believe that the linear rank p-value approximations are adequate when consideration is restricted to sample size and percent censoring.

Log-linear analyses: All p-values derived from log-linear analyses are based on large sample chi-square approximations. The adequacy of these approximations has generally been studied in terms of the magnitudes of the expected cell counts in multiway contingency tables. There is extensive literature on this subject with resultant guidance published in recent statistical texts. Conover (15) states that the chi-square approximation is good if the expected cell counts are fairly large but if some of the expected counts are small, the approximation may be poor. He quotes Cochran (16), who concluded that, if any of the expected counts are less than 1 or if more than 20% are less than 5, the approximation may be poor. Conover views

Cochran's advice as, perhaps, too conservative and renders the opinion that the expected counts may be as small as 1 without endangering the validity of the test. Since most expected counts in this report are greater than 5, the chi-square approximations are considered adequate by the authors of this report.

SMR analyses: Large sample chi-square approximations were used to obtain the p-values in the SMR analyses. The first of these was for a likelihood ratio test for the hypothesis that the data satisfies the product model (7), the second was for a likelihood ratio test that the SMR was equal to unity. The test of fit for the product model is analogous to a test for no three-factor interaction in a log-linear model, the factors being survival states (dead, alive), group (Ranch Hand, Comparison) and year-of-birth. Sample size requirements for this procedure are, therefore, the same as those described above for log-linear analysis; that is, that the expected numbers of dead at each level of year-of-birth be at least 5 or at least 1, depending on the advice of Conover and Cochran. The test for an SMR equal to unity is not analogous to a test on the main effect in the same log-linear model. No guidance has been published regarding the sample size requirements for the adequacy of the chi-square approximation. In our opinion, this approximation is adequate in these data.

Gail and Ware analysis: The test statistic for comparing an observed survival distribution with a reference life table is a standardized sum of deviations between observed and expected numbers of deaths and has, for large samples, an approximate standard normal distribution under the null hypothesis. The minimum sample size and maximum censoring percentage needed for this approximation to be adequate is not known. In our opinion, this approximation is adequate in these data.

Ejigou-McHugh relative risk analyses: The statistic used in testing relative risk equal to unity has an approximate standard normal distribution under the null hypothesis when the number of matched sets is large. In-house simulations have shown that this approximation is adequate with 1200 match sets. The threshold of adequacy has not been investigated to date. In our judgement, the approximation is good in these analyses.

7.2 Assumptions and Statistical Assessment of their Validity

In all studies, statistical procedures are based upon assumptions regarding the data. Good statistical practice requires that the assumptions be checked before proceeding to the final analysis. In most cases this is done subjectively by examining plots of the data. For some statistical procedures, the assumptions can be tested directly; such tests are termed pretests. When resampling is not possible, pretesting should be accounted for in the overall inference. Unfortunately, pretests and procedures which account for pretests in the overall inference are almost nonexistent in the field of statistics. Of the five procedures used in this report, a pretest of assumptions exists only for the SMR analysis, and it is not currently known, how to take that pretesting into consideration in the overall analysis. Generally, pretesting should be carried out so that the overall significance level of the pretests and the final inferential test should be a prescribed value, such as 5%.

Linear rank tests: The logrank and Wilcoxon procedures are based upon the assumptions that the underlying survival distributions are continuous, that survival and censoring are statistically independent and that the difference in group sur-

vival distributions does not change with levels of the stratification variable. The third assumption would hold if, for example, there were no interaction terms involving group membership in the accelerated failure time model (5). In our opinion, the first two of these assumptions can be safely assumed in these analyses. The third must be checked. There does not exist a statistical procedure for testing the assumption that the difference in group survival distributions does not change with levels of the stratification variable, without making further assumptions. If further assumptions were made and such a test were developed, there would, at this time, be no way to adjust its critical value so that the overall significance level was 5%. In this report, the stratification variables were year-of-birth, race, and occupation. The no-interaction assumption was subjectively checked by comparing the logrank and Wilcoxon values with other analyses, looking for consistency. There is some indication that the assumption is not met in the officer subgroup and, therefore, the logrank and Wilcoxon values are misleading for contrasting Ranch Hand and Comparison officers.

Log-linear analyses: The log-linear analyses are based upon the assumptions that the data are distributed as multinomials or product-multinomials, that all interactions of order higher than the one of interest are nonexistent and that there is no confounding. The multinomial assumption is correct in these analyses because the data were categorized so that the multinomial or product-multinomial model would hold. Tests for the existence of interactions of all orders are available and are carried out in all analyses but, at this time, there is no way to adjust their critical values so that the significance level of the overall procedure is 5%. Statisticians typically use a 5% significance level for each pretest, but this may vary.

SMR analyses: The basic assumption in these analyses is that relative risk is constant across levels of the stratification variable. In these analyses the stratification variable is year-of-birth. A likelihood ratio test was used to check this assumption. It is not known how to prescribe its critical value so that the overall level of significance is 5%. This assumption was also checked using additional log-linear analyses.

Gail and Ware analyses: The basic assumption in these analyses is that the study hazard function is proportional to the reference hazard function. There does not exist a single sample test for the proportional hazards assumption. This assumption was checked subjectively by computing relative risks at different ages within the data sets.

Ejigou-McHugh relative risk analyses: This analysis assumes that relative risk is constant with respect to the matching variables. A procedure for testing this assumption has been recently developed (17) but has not yet been programmed for inclusion in these reports. The new method does not provide for the adjustment of the pretest critical value so that the overall significance level is 5%. This assumption was subjectively checked in this report by comparing the Ejigou-McHugh relative risk with the SMR, looking for consistency.

7.3 Summary

The issues regarding large sample approximations and pretesting assumptions are intrinsic to the field of mathematical statistics and, therefore, are relevant to

applications of statistical theory in any research. In this respect, the statistical content of this report reflects the extent of current theory.

8. Future Commitments

Future work will attempt to evaluate mortality patterns as a function of occupational subgroup in the ground cohort. This effort will require the collection of data to delineate differential exposure between occupational subgroups. Flight line duties and herbicide contact will be ascertained objectively, along with additional medical risk factors, occupational exposures and socioeconomic factors. These analyses will be increasingly meaningful as the population ages and mortality rates permit use of more incisive statistical tools. Joint morbidity-mortality analyses, adjusting for relevant covariates will be carried out. Finally, the small sample properties of the linear rank, relative risk, and SMR tests will be investigated by simulation and analytical methods.

9. Summary and Conclusion

Evaluation of summary counts of death by rank and occupation did not reveal any statistically significant differences between the Ranch Hand and Comparison groups. Other mortality analyses described in this report have revealed some differences in death experience between the herbicide/dioxin exposed group, their matched Comparisons and other external Comparison groups.

Overall mortality of the Ranch Hand group (4.4%) is nearly identical to that of Ranch Hand officers have experienced fewer deaths the Comparison group (4.6%). than the Comparison group officers, but this difference is not statistically signi-There is an interaction in these data, however. Ranch Hand officers born between 1905 and 1935 have experienced fewer deaths than Comparison officers born during the same era. On the other hand, Ranch Hand officers born after 1935 have experienced more deaths than their Comparisons. Although these differences within birth-year strata are not statistically significant, this change in the group by survival status relationship with birth year is statistically significant. tionally, Ranch Hand officers experienced fewer deaths after age 35 years than did Comparison officers, while Ranch Hand officers experienced more deaths before age 35 years than did Comparisons. Further research will investigate whether there is any association between birth year and age of death and mortality patterns in these officer cohorts.

At this time, Ranch Hand ground and enlisted personnel have experienced more nortality and Ranch Hand flying personnel have experienced lower mortality than their Comparisons, but these differences are not statistically significant. Preliminary analyses using exposure indices have indicated no association between herbicide exposure in either the officer, enlisted, flying or ground Ranch Hand subgroups.

Both Ranch Hand and Comparison officers have experienced less mortality than Ranch Hand or Comparison enlisted personnel. Ranch Hand flying personnel have experienced less mortality than Ranch Hand ground personnel, while Comparison flying and ground personnel have experienced similar mortality patterns.

Examining causes of death, Ranch Hand officer and flying groups have experienced fewer deaths from cardiovascular disease and cancer than have the Comparisons, but this difference is not statistically significant. No apparent specific disease excesses were noted in the Ranch Hand ground or enlisted groups relative to their Comparisons. All Ranch Hand cohorts are elevated in the category of digestive system deaths, but this difference is not statistically significant. There was a single case of soft tissue sarcoma in the Comparison group, and no cases occurred in the Ranch Handers.

The Ranch Hand and Comparison groups were contrasted with five external populations. All study groups are experiencing significantly less mortality than U.S. White males. All study groups except Ranch Hand enlisted personnel are experiencing statistically significantly less mortality than the corresponding nondisability retired DOD population. The Ranch Hand enlisted mortality is not significantly different from the nondisability retired DOD enlisted population. The Ranch Hand and Comparison groups taken together have experienced a mortality pattern not statistically different from civil service employees.

In conclusion, summary counts of death by rank and occupation did not reveal any statistically significant differences, within the power limitations of this study, between the Ranch Hand and Comparison groups. This study has excellent power of detecting a doubling of risk of death, and therefore it is unlikely that an effect of this magnitude could have been missed. Ranch Hand officers born between 1905 and 1935 have experienced favorable mortality relative to their Comparisons while the converse is true for officers born after 1935. Analogous patterns are seen in officers, conditioned on age at death. Although Ranch Hand ground personnel experienced less favorable mortality relative to Comparisons irrespective of date of birth or age at death, this difference is not statistically significant. index analyses indicate these mortality rate differences cannot be attributed to herbicide exposure. These analyses have identified several findings of interest, which will be further evaluated in future mortality updates. The findings of this report are similar to those of prior mortality analyses with the exception that the non-Black Ranch Hand enlisted personnel now demonstrate statistically significantly better survival than the 1978 U.S. White male population.

References

- Lathrop, G. D., Moynahan, P. M., Albanese, R. A., Wolfe, W. H. (1983).
 An Epidemiologic Investigation of Health Effects in Air Force Personnel Following Exposure to Herbicides: Baseline Mortality Study Results. (NTIS Order Number: AD-A130 793)
- Lathrop, G. D., Wolfe, W. H., Albanese, R. A., Moynahan, P. M. (1984).
 An Epidemiologic Investigation of Health Effects in Air Force Personnel Following Exposure to Herbicides: Baseline Morbidity Study Results. (NTIS Order Number: AD-A138 340)
- Wolfe, W. H., Michalek, J. E., Albanese, R. A., Lathrop, G. D., Moynahan, P. M. (1984). An Epidemiologic Investigation of Health Effects in Air Force Personnel Following Exposure to Herbicides: Mortality Update - 1984. (NTIS Order Number: not yet available)
- 4. Kaplan, E. L. and Meier, P. (1958). Nonparametric estimation from incomplete observation. <u>Journal of the American Statistical</u>
 Association 53:457-481.
- 5. Prentice, R. L., (1978). Linear rank tests with right censored data. Biometrika 65:167-179.
- 6. Ejigou, A. and McHugh, R. (1981). Relative risk estimation under multiple matching. Biometrika 68:85-91.
- 7. Gail, M. (1978). The analysis of heterogeneity for indirect standardized mortality ratios. <u>Journal of the Royal Statistical</u> Society, A, 141:224-234.
- 8. Vital Statistics of the United States, 1978 Vol II Section 5, Life Tables; US Dept of Health and Human Services, DHHS Publication No (PHS) 81-1104; Hyattsville, Maryland; 1980.
- 9. Evaluation of the Military Retirement System FY 1980. Office of the Actuary, Defense Manpower Data Center, 300 North Washington Street, Alexandria, Virginia 22314.
- 10. Board of Actuaries of the Civil Service Retirement System, Fifty-Seventh Annual Report. US Government Printing Offices; 1980.
- 11. Bishop, Y. M., Fienberg, S. E., and Holland, P. W. (1975). <u>Discrete Multivariate Analysis: Theory and Practice</u>. The MIT Press, Cambridge.
- 12. Gail, M. and Ware, J. H. (1979). Comparing observed life table data with a known survival curve in the presence of random censorship.

 Biometrics 35, 385-391.

- 13. Latta, R. B. (1981). A Monte Carlo study of some two sample rank tests with censored data. <u>Journal of the American Statistical Association</u>, 76, 713-718.
- 14. Michalek, J. E., Mihalko, D. and White, T. (1985). A Monte Carlo study of logrank, Wilcoxon and normal scores procedures on matched and censored data. To appear in Communications in Statistics.
- 15. Conover, N. J. (1980) <u>Practical Nonparametric Statistics</u>, John Wiley, New York.
- 16. Cochran, W. G. (1952). The chi-square test of goodness of fit. Annals of Mathematical Statistics, 23, 315-345.
- 17. Ejigou, A. and McHugh, R. (1984). Testing the homogeneity of relative risk under multiple matching. Biometrika, 71, 408-411.

PRINCIPAL INVESTIGATORS

William H. Wolfe, MD, MPH, FACPM Colonel, USAF, MC Chief, Epidemiology Division

Joel E. Michalek, PhD, GS-14 Research Mathematical Statistician Data Sciences Division

CONTRIBUTORS

Richard A. Albanese, MD, GM-15 Chief, Biomathematical Modeling Branch Data Sciences Division

Vincent V. Elequin, BS, RRA, GS-11 Medical Record Librarian Epidemiology Division

Melody Prihoda, MS, GS-9 Mathematical Statistician Data Sciences Division

Alton J. Rahe, MS, GS-13 Mathematical Statistician Data Sciences Division

Thomas J. White, MA Senior Subject Matter Specialist Data Sciences Division

Appendix Table 1

Ranch Hand Officers Versus Ranch Hand Enlisted Mortality by Year-Of-Birth (SMR = 0.515, P1 = 0.27, P2 = 0.047)

	Ranch H	land Of		Ranch Hand Enlisted		
Birth <u>Year</u>	At Risk	Dead	Rate per 100	At Risk	<u>Dead</u>	Rate per 100
1905-1924	41	3	7.3	29	7	24.1
1925-1934	194	5	2.6	195	15	7.7
1935-1939	9 5	4	4.2	116	3	2.6
1940-1944	91	2	2.2	119	3	2.5
1945-1954	45	2	4.4	332		3.3
Total	466	16		791	39	

Appendix Table 2

Comparison Officers Versus Comparison Enlisted Mortality by Year-Of-Birth (SMR = 0.648, P1 = 0.88, P2 = 0.001)

	Officers			Enlisted		
Birth <u>Year</u>	At Risk	Dead	Rate per 100	At Risk	Dead	Rate per 100
1905-1919 1920-1924 1925-1929 1930-1934 1935-1939 1940-1944 1945-1954	44 161 290 640 458 495	4 17 20 32 13 7	9.1 10.6 6.9 5.0 2.8 1.4 2.6	66 80 211 749 562 601 1624	13 13 26 47 26 17 45	19.7 16.2 12.3 6.3 4.6 2.8 2.8
Total	2278	98		3893	187	

Appendix Table 3

Ranch Hand Flying Personnel Versus Ranch Hand Ground Personnel Mortality by Year-Of-Birth (SMR = 0.572, P1 = 0.41, P2 = 0.067)

	Flyers				Ground		
Birth		-	Rate			Rate	
Year	At Risk	Dead	<u>per 100</u>	At Risk	<u>Dead</u>	<u>per 100</u>	
1905-1924	44	4	9.1	26	6	23.1	
1925-1934	272	10	3.7	117	10	8.5	
1935-1939	145	6	4.1	66	1	1.5	
1940-1944	121	2	1.7	89	3	3.4	
1945-1954	_64	_2	3.1	313	11	3.5	
Total	646	24		611	31		

Appendix Table 4

Comparison Flying Versus Comparison Ground Personnel Mortality by Year-Of-Birth
Within Comparison Group
(SMR = 0.909 P1 = 0.46, P2 = 0.65)

Flyers Ground Birth Rate Rate per 100 Year At Risk Dead per 100 At Risk Dead 1905-1919 16.9 45 6 13.3 65 11 15.2 1920-1924 175 20 11.4 66 10 . 1925-1929 13.2 350 26 7.4 151 20 5.0 1930-1934 6.0 21 966 58 423 3.7 1935-1939 4.0 698 26 322 13 1940-1944 653 15 2.3 443 9 2.0 1945-1954 40 2.6 276 10 3.6 1538 Total 3163 161 3008 124